

UNIFORM LIMIT THEOREMS FOR THE MAXIMUM CUMULATIVE SUM IN PROBABILITY⁽¹⁾

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1. **Introduction.** Let $\{x_k : k=1, 2, 3, \dots\}$ be a sequence of independent random variables with mean zero, variance one, and for example bounded absolute third moments. Set $s_n^+ = \max_{1 \leq k \leq n} s_k$ for $s_k = x_1 + x_2 + \dots + x_k$ and choose $\lambda \geq 0$. Then, it was shown by Erdős and Kac [5] in 1946 that

$$\lim_{N \rightarrow \infty} P(s_N^+ < \lambda(N)^{1/2}) = M(\lambda) = (2/\pi)^{1/2} \int_0^\lambda e^{-u^2/2} du,$$

where $M(\lambda) = 0$ for $\lambda < 0$. The method was an ingenious use of the multi-dimensional central limit theorem; Kai Lai Chung [3] estimated the errors involved to calculate

$$(1.1) \quad \sup_\lambda |P(s_N^+ < \lambda(N)^{1/2}) - M(\lambda)| = O\left(\frac{\log N}{N^{1/26}}\right).$$

In general, if $\{x_k : k=1, 2, 3, \dots\}$ is an arbitrary sequence of independent random variables with means zero, let $q_N^2 = E(x_1^2) + E(x_2^2) + \dots + E(x_N^2)$ and $\{x_t : 0 \leq t \leq 1\}$ be Brownian motion. In §2 we show

THEOREM 1. *For all $p > 2$ and all N ,*

$$(1.2) \quad \sup_\lambda |P(s_N^+ < \lambda q_N) - M(\lambda)| \leq 300 \left(\frac{1}{q_N^p} \sum_{k=1}^N E(|x_k|^p) \right)^{1/p+1}.$$

THEOREM 2. *For all $p > 2$ and all N ,*

$$(1.3) \quad \sup_\alpha \sup_\beta \left| P(s_N^+ \leq \alpha q_N, s_N \leq \beta q_N) - P\left(\max_{0 \leq t \leq 1} x_t \leq \alpha, x_1 \leq \beta\right) \right| \\ \leq 800 \left(\frac{1}{q_N^p} \sum_{k=1}^N E(|x_k|^p) \right)^{(2/3)/[p + (2/3)]}.$$

For example, if $E(x_k^2) \equiv 1$ and $E(|x_k|^p) \leq C$,

$$(1.4) \quad \sup_\lambda |P(s_N^+ < \lambda(N)^{1/2}) - M(\lambda)| = O\left(\frac{1}{N^{(p-2)/(2p+2)}}\right)$$

or $O(1/n^{1/8})$ under the hypotheses of (1.1). Similarly, if $N = O(q_N^2)$ and

$$\sup_k P(|x_k| > \lambda) \leq C e^{-\beta \lambda^\epsilon}$$

for $\beta > 0, \epsilon > 0$, then by arguing as in [11]

$$(1.5) \quad \sup_\lambda |P(s_N^+ < \lambda q_N) - M(\lambda)| = O\left(\frac{(\log N)^{1/\epsilon}}{(N)^{1/2}}\right).$$

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The limit distribution in (1.3) is of course a result of the Invariance Principle ([4], [9]), and is essentially equivalent (see §2) to its special case

$$\sup_{\alpha \geq 0} \sup_{\beta \geq 0} |P(s_N^+ > \alpha q_N, s_N \leq (\alpha - \beta)q_N) - \Phi^c(\alpha + \beta)| \\ \leq 1600 \left(\frac{1}{q_N^p} \sum_{k=1}^N E(|x_k|^p) \right)^{(2/3)/[p + (2/3)]}$$

where $\Phi(\lambda) = (1/(2\pi)^{1/2}) \int_{-\infty}^{\lambda} e^{-u^2/2} du$ and $\Phi^c(\lambda) = 1 - \Phi(\lambda)$. A uniform rate of convergence here was first obtained by Hirsch [7]⁽²⁾, with, relative to (1.3), restrictions on $\{x_k\}$, α , β , and a weaker rate of convergence. This was used to prove a theorem of iterated logarithm type. In particular, the limit theorem of Hirsch can be extended to classes of $\{x_k\}$ without third moments or essentially constant variances.

As another example, consider the random walk on the real line given by $s_0 = 0$, $s_N = x_1 + x_2 + \cdots + x_N$. The probability of moving initially to the left, and of remaining in the left half line for at least N time units, can be estimated by the right hand side of (1.2). In particular, if $|x_k| \leq C$ a.s., this probability is bounded by $300 C/q_N$. This last result seems to be due initially to Burkholder⁽³⁾, by the use of martingale methods.

Added in proof. Burkholder's approach apparently works only if $x_k = \pm a_k$ a.s. for constants $\{a_k\}$, $|a_k| \leq C$, and uses Theorem 8 in [2]. Theorem 1 would then give a generalization.

For $p = 3$, Theorem 1 can be obtained within a logarithmic factor, with the same bound for (1.2) and (1.3), by the methods of [9, §5]. For q_N^2 essentially a constant times N , the estimate (1.2) (multiplied by $(\log N)^{1/2}$) follows by arguing as in [10] for $2 < p < 4$ or as in [11] for $5 < p < \infty$. However none of these arguments are elementary; it is one of the purposes here to show that these results can be obtained from the Reflection Principle of Désirée André and Lévy.

Finally, it should be mentioned that no attempt was made to find the optimal constants in (1.2) and (1.3). If $|x_k| \leq C$ and $\sum_{k=1}^N E(x_k^2) = 1$, it follows from (1.2) or Lemma 3, §2 that

$$(1.6) \quad \sup_{\lambda} |P(s_N^+ < \lambda) - M(\lambda)| \leq 300 C/(N)^{1/2}.$$

By [1] or e.g. $x_k = \pm N^{-1/2}$, this is the best possible rate of convergence. Whether or not (1.2), or in particular (1.5), is best possible seems to be an open question.

2. Proofs.

LEMMA 1 (CENTERING). *Let x_1, x_2, \dots, x_N be independent random variables such that*

$$(2.1) \quad E(x_k) = 0, E(|x_k|^3) \leq \Omega E(x_k^2), \quad 1 \leq k \leq N.$$

Then, if $s_N = x_1 + x_2 + \cdots + x_N$ and $\mu = 50\Omega$,

$$(2.2) \quad P(s_N \leq -\mu) < \frac{1}{2}.$$

⁽²⁾ See also [1] for theorems of Cramér type.

⁽³⁾ Colloquium talk 1-9-67, discussing results in [2].

Proof. By the Berry-Esseen theorem [6, p. 521]

$$\begin{aligned} \left| P(s_N \leq \lambda) - \Phi\left(\frac{\lambda}{q_N}\right) \right| &\leq \frac{33}{4} \frac{\Omega}{q_N}, \quad \text{all } \lambda, \\ P(s_N \leq -\mu) &\leq \Phi\left(-\frac{\mu}{q_N}\right) + \frac{33}{4} \frac{\Omega}{q_N} \\ &= \frac{1}{2} + \frac{33}{4} \frac{\Omega}{q_N} - \frac{1}{(2\pi)^{1/2}} \int_0^{\mu/q_N} e^{-u^2/2} du. \end{aligned}$$

Now, if $0 \leq q_N < \mu/(2)^{1/2}$, then $P(s_N \leq -\mu) \leq (1/\mu^2)q_N^2 < \frac{1}{2}$ by Tchebysheff's inequality. But if $q_N \geq \mu/(2)^{1/2}$, by the above

$$P(s_N \leq -\mu) \leq \frac{1}{2} + \frac{\Omega}{q_N} \left(\frac{33}{4} - \frac{50}{(2\pi)^{1/2}} \int_0^1 e^{-u^2} du \right) < \frac{1}{2}.$$

LEMMA 2 (REFLECTION). Let x_1, x_2, \dots, x_N be independent random variables, such that for a fixed $\mu \geq 0$ and $1 \leq k \leq N$

$$(2.3) \quad P(s_N - s_k \geq \mu) \leq \frac{1}{2}, \quad P(s_N - s_k \leq -\mu) \leq \frac{1}{2}.$$

Then, for all $\varepsilon > 0$, $\lambda > 0$,

$$(2.4) \quad \begin{aligned} 2P(s_N > \lambda + \varepsilon + \mu) - \sum_1^N P(x_k > \varepsilon) &\leq P(s_N^+ > \lambda) \\ &\leq 2P(s_N > \lambda - \mu). \end{aligned}$$

Proof. This is standard; see e.g. [8, p. 247, p. 558] or argue as in Lemma 5.

LEMMA 3. Let x_1, x_2, \dots, x_N be independent random variables such that

$$(2.5) \quad E(x_k) = 0, \quad |x_k| \leq C \text{ a.s.}, \quad \sum_1^N E(x_k^2) = 1$$

for some constant C and $1 \leq k \leq N$. Then, for all λ ,

$$(2.6) \quad |P(s_N^+ < \lambda) - M(\lambda)| < 60C.$$

Proof. Clearly (2.5) implies (2.1) for $\Omega = C$, and hence (2.3) for $\mu = 50C$. Also, $|P(s_N > \lambda) - \Phi^c(\lambda)| \leq 33C/4$ by Berry-Esseen. Hence by (2.4) with $\varepsilon = C$, $\lambda \geq 0$

$$\begin{aligned} 2\Phi^c(\lambda + \varepsilon + \mu) - \frac{33}{2}C &\leq P(s_N^+ > \lambda) \leq 2\Phi^c(\lambda - \mu) + \frac{33}{2}C, \\ |P(s_N^+ > \lambda) - 2\Phi^c(\lambda)| &\leq \frac{33}{2}C + (2/\pi)^{1/2}(\mu + C) < 60C. \end{aligned}$$

Proof of Theorem 1. First, note that it is sufficient to assume $q_N = 1$ and $C = (\sum_1^N E(|x_k|^p))^{1/(p+1)} < \frac{1}{2}$. Define $y_k = x_k$ if $|x_k| \leq C$; otherwise set $y_k = 0$. Then $q_c^2 = \sum_1^N \sigma^2(y_k) \geq \frac{1}{4}$, and $|y_k - E(y_k)|/q_c \leq 4C$. Applying Lemma 3 to the $\{y_k\}$, and in general arguing as in [11, §5], we arrive at (1.2).

REMARK. If one does not object to an extra condition of the form $\max_{1 \leq k \leq n} E(|x_k|^3)/E(x_k^2) \leq Cq_n^{1-\theta}$, then (1.2) can be obtained directly from Lemma 1 by setting

$$\varepsilon = \left(q_N \sum_1^N E(|x_k|^p) \right)^{1/(p+1)} \text{ in (2.4).}$$

LEMMA 4. Let x_1, x_2, \dots, x_N be as in (2.5), and for some fixed $\lambda > 0$ set $\tau = \min \{k : s_k > \lambda\}$ if $s_N^+ > \lambda$ and set $\tau = \infty$ if $s_N^+ \leq \lambda$. Assume $1 \leq k \leq N$ and $\frac{1}{2} \leq q_k \leq q_N = 1$. Then

$$(2.7) \quad P(k < \tau \leq N) \leq 1 - q_k + 200C.$$

Proof. We conclude (2.3) for $\mu = 50C$ as before, and by (2.4)

$$\begin{aligned} P(k < \tau \leq N) &= P(s_N^+ > \lambda) - P(s_k^+ > \lambda) \\ &\leq 2\Phi^c(\lambda - \mu) - 2\Phi^c\left(\frac{\lambda + \mu + C}{q_k}\right) + \frac{33}{2} C \left(1 + \frac{1}{q_k}\right) \\ &\leq (2/\pi e)^{1/2} \frac{1 - q_k}{q_k} + 200C, \text{ etc.} \end{aligned}$$

LEMMA 5. For x_1, x_2, \dots, x_N as in Lemma 4,

$$(2.8) \quad \sup_{\alpha \geq 0} \sup_{\beta \geq 0} |P(s_N^+ > \alpha, s_N \leq \alpha - \beta) - \Phi^c(\alpha + \beta)| \geq 200C^{2/3}.$$

Proof. First, by the Berry-Esseen theorem, for all $\beta \geq 0, \alpha \geq 0$,

$$P(s_N - s_k \leq -\beta) \leq P(s_N - s_k \geq \beta) + \frac{33}{2} \frac{C}{(1 - q_k^2)^{1/2}},$$

$$\begin{aligned} P(s_N^+ > \alpha, s_N \leq \alpha - \beta) &\leq \sum_1^{N-1} P(s_1 \leq \alpha, \dots, s_{k-1} \leq \alpha, s_k > \alpha) P(s_N - s_k \leq -\beta) \\ &\leq \sum_1^{N-1} P(s_1 \leq \alpha, \dots, s_{k-1} \leq \alpha, s_k > \alpha) P(s_N - s_k \geq \beta) + Q \\ &\leq P(s_N > \alpha + \beta) + Q. \end{aligned}$$

Here

$$Q = \frac{33}{2} \frac{C}{(1 - q_l^2)^{1/2}} + P(l < \tau \leq N) \leq \frac{33}{2} \frac{C}{(1 - q_l)^{1/2}} + 1 - q_l + 200C$$

for some l to be determined, $q_l \geq \frac{1}{2}$. Similarly, since $|x_k| \leq C$,

$$\begin{aligned} P(s_N^+ > \alpha, s_N \leq \alpha - \beta) &\geq \sum_1^{N-1} P(s_1 \leq \alpha, \dots, s_{k-1} \leq \alpha, s_k > \alpha) P(s_N - s_k \leq -\beta - C) \\ &\geq \sum_1^{N-1} P(s_1 \leq \alpha, \dots, s_{k-1} \leq \alpha, s_k > \alpha) P(s_N - s_k \geq \beta + C) - Q \\ &\geq P(s_N > \alpha + \beta + 2C) - Q, \end{aligned}$$

and

$$|P(s_N^+ > \alpha, s_N \leq \alpha - \beta) - \Phi^c(\alpha + \beta)| \leq \frac{33}{4} C + \frac{2}{(2\pi)^{1/2}} C + Q.$$

Note that unless $C < 1/8$, (2.8) is trivial. If now $k = \max \{j : q_j < 1 - C^{2/3}\}$, then $C^{2/3} < 1 - q_k < C^{2/3} + C^2 \leq 2C^{2/3} < \frac{1}{2}$. Thus $Q < 180C^{2/3}$ and (2.8).

Proof of Theorem 2. I claim that under the hypotheses of Lemma 5, for all α and all β

$$(2.9) \quad \left| P(s_N^+ \leq \alpha, s_N \leq \beta) - P\left(\max_{0 \leq t \leq 1} x_t \leq \alpha, x_1 \leq \beta\right) \right| \leq 210C^{2/3}.$$

Clearly we can assume $C < 1$. Now, if $\alpha < 0$, then $P(s_N^+ \leq \alpha, s_N \leq \beta) \leq P(s_N^+ \leq 0) < 60C < 210C^{2/3}$, while $P(\max_{0 \leq t \leq 1} x_t \leq 0) = 0$. Also if $0 \leq \alpha \leq \beta$,

$$P(s_N^+ \leq \alpha, s_N \leq \beta) = P(s_N^+ \leq \alpha) \leq M(\alpha) + 60C,$$

and since $P(\max_{0 \leq t \leq 1} x_t \leq \alpha) = M(\alpha)$ [6, p. 171], (2.9) follows from Lemma 3. Finally, if $\alpha \geq 0, \beta \leq \alpha$, then

$$P(s_N^+ \leq \alpha, s_N \leq \beta) = P(s_N \leq \beta) - P(s_N^+ > \alpha, s_N \leq \beta),$$

and (2.9) follows from (2.8), the Berry-Esseen theorem, and the reflection principle for Brownian motion. Finally, given (2.9) for the hypotheses (2.5), Theorem 2 follows by a truncation argument as in Theorem 1 or [11, §5].

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